# Probabilistic Convergence and Bounds

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# **Useful Inequalities**

• Markov's Inequality:

Non-negative r.v.  $Z \ge 0$  and real a > 0

 $\Pr(Z \ge a) \le \frac{\mathbb{E}[Z]}{a}$ 

• Chebyshev's Inequality:

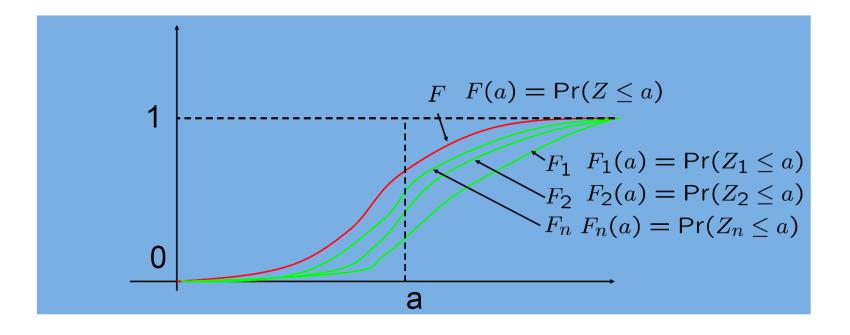
R.v. X with  $\mathbb{E}[X] < \infty$ , real a > 0

 $\Pr(|X - \mathbb{E}[X]| \ge a) \le \frac{\operatorname{Var}(X)}{a^2} \qquad (\operatorname{Set} Z \equiv |X - \mathbb{E}[X]|^2.)$ 

#### **Convergence in Distribution**

- Notation:  $Z_n \xrightarrow{d} Z, Z_n \xrightarrow{D} Z$
- **Definition:** Let  $F_n$ , F be cdfs of  $Z_n$ , and Z:

 $\lim_{n \to \infty} F_n(\omega) = F(\omega) \,\,\forall \,\, \omega \in \mathbb{R} \text{ s.t. } F \text{ is continuous}$ 



# Convergence in P<sup>th</sup> Mean, L<sub>p</sub> norm

- Notation:  $Z_n \stackrel{L_p}{\to} Z$ . If  $p = 2, Z_n \stackrel{q.m.}{\to} Z$  (in quadratic mean)
- **Definition:** For fixed  $p \ge 1$ ,  $\lim_{n \to \infty} \mathbb{E}[|Z_n(\omega) Z(\omega)|^p] = 0$
- Intuition:

Note that for fixed n,  $\mathbb{E}[|Z_n(\omega) - Z(\omega)|^p]$  is a deterministic value. Let  $a_n \equiv \mathbb{E}[|Z_n(\omega) - Z(\omega)|^p]$ Hence,  $a_n \to 0$ 

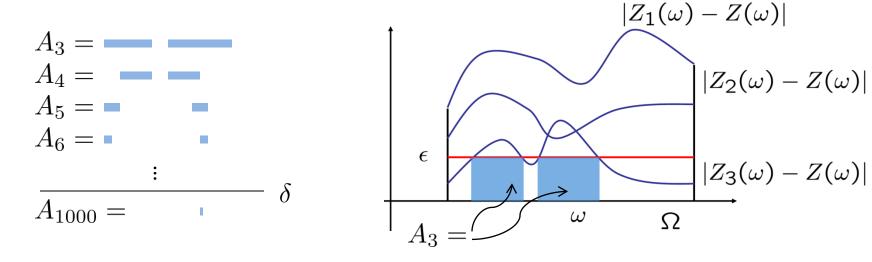
#### **Almost Sure Convergence**

- Notation:  $Z_n \stackrel{a.s.}{\rightarrow} Z, Z_n \to Z \text{ w.p. } 1$
- **Definition:**  $P\left(\omega \in \Omega : \lim_{n \to \infty} Z_n(\omega) = Z(\omega)\right) = 1$
- Intuition:

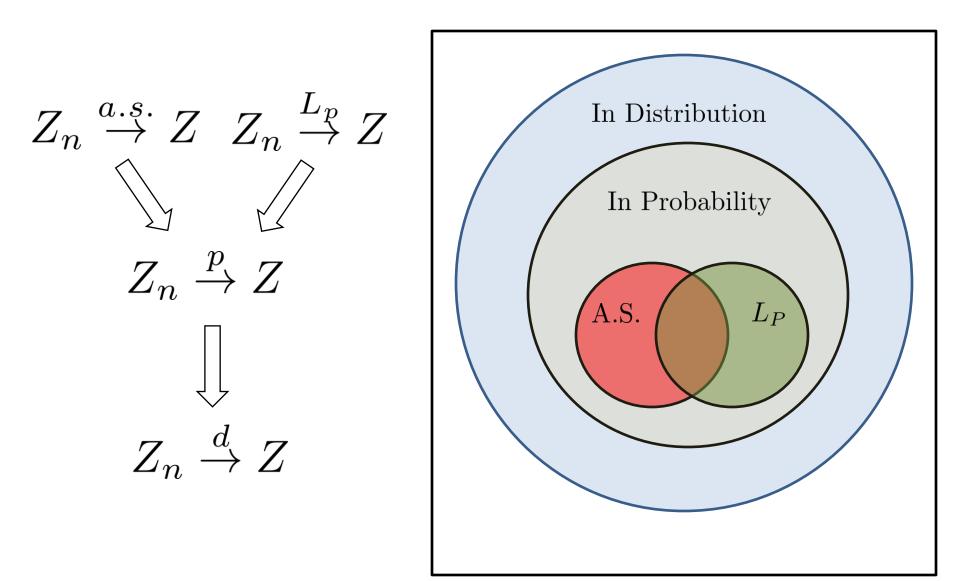
Let 
$$A \equiv \left\{ \omega \in \Omega : \lim_{n \to \infty} Z_n(\omega) = Z(\omega) \right\}$$
, then  $P(A) = 1$ .  
E.g.  $\Omega = [0, 1], P = \text{Unif}[0, 1]$ .  
 $A = \underbrace{0}_{0} \underbrace{0}_{1}$ 

#### **Convergence in Probability**

- Notation:  $Z_n \xrightarrow{p} Z$
- **Definition:**  $\forall \epsilon > 0 \lim_{n \to \infty} P(\omega \in \Omega : |Z_n(\omega) Z(\omega)| \ge \epsilon) = 0$  $\forall \epsilon > 0 \lim_{n \to \infty} P(|Z_n - Z| \ge \epsilon) = 0$
- Intuition:  $\forall \epsilon > 0 \ \forall \delta > 0 \ \exists N \text{ s.t. } n > N \implies P(|Z_n Z| \ge \epsilon) < \delta$ Let  $\epsilon > 0, \ \delta > 0$  be given Define  $A_n \equiv \{\omega \in \Omega : |Z_n(\omega) - Z(\omega)| \ge \epsilon\}$



#### **Relation Among Convergences**



Suppose  $Z_n \stackrel{q.m.}{\to} Z$ ; i.e.,  $\lim_{n \to \infty} \mathbb{E}[|Z_n - Z|^2] = 0.$ 

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Let  $\epsilon > 0$ , and  $\delta > 0$  be given.

Let N be s.t.  $n > N \implies \mathbb{E}[|Z_n - Z|^2] \le \epsilon^2 \delta$ 

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Let  $\epsilon > 0$ , and  $\delta > 0$  be given.

Let N be s.t.  $n > N \implies \mathbb{E}[|Z_n - Z|^2] \le \epsilon^2 \delta$ 

Then, by Markov's:  $\Pr(|Z_n - Z|^2 > \epsilon^2) \leq \frac{\mathbb{E}[|Z_n - Z|^2]}{\epsilon^2}$ 

Suppose  $Z_n \stackrel{q.m.}{\to} Z$ ; i.e.,  $\lim_{n \to \infty} \mathbb{E}[|Z_n - Z|^2] = 0.$ 

Let  $\epsilon > 0$ , and  $\delta > 0$  be given.

Let N be s.t.  $n > N \implies \mathbb{E}[|Z_n - Z|^2] \le \epsilon^2 \delta$ 

Then, by Markov's:  $\Pr(|Z_n - Z|^2 > \epsilon^2) \leq \frac{\mathbb{E}[|Z_n - Z|^2]}{\epsilon^2}$ 

So,  $n > N \implies \Pr(|Z_n - Z|^2 > \epsilon^2) \le \frac{\mathbb{E}[|Z_n - Z|^2]}{\epsilon^2} \le \frac{\epsilon^2 \delta}{\epsilon^2} = \delta$ 

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So,  $n > N \implies \Pr(|Z_n - Z|^2 > \epsilon^2) \le \frac{\mathbb{E}[|Z_n - Z|^2]}{\epsilon^2} \le \frac{\epsilon^2 \delta}{\epsilon^2} = \delta$ 

Hence,  $\forall \epsilon > 0 \ \forall \delta > 0 \ \exists N \text{ s.t. } n > N \implies P(|Z_n - Z| \ge \epsilon) < \delta.$ 

Suppose  $Z_n \xrightarrow{p} Z$ ; i.e.,  $\forall \epsilon > 0 \lim_{n \to \infty} P(|Z_n - Z| \ge \epsilon) = 0$ 

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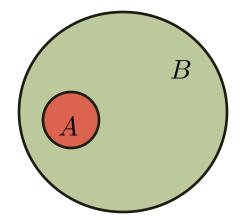
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=  $P(Z_n \le z, Z \le z + \epsilon) + P(Z_n \le z, Z > z + \epsilon)$ 

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=  $P(Z_n \le z, Z \le z + \epsilon) + P(Z_n \le z, Z > z + \epsilon)$   
 $\le P(Z \le z + \epsilon) + P(Z > Z_n + \epsilon)$ 

Since, If  $A \implies B$ , then  $\Pr(A) \leq \Pr(B)$ .



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 $\le F(z + \epsilon) + P(Z - Z_n > \epsilon) + P(Z - Z_n < -\epsilon)$ 

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$$F(z - \epsilon) = P(Z \le z - \epsilon)$$
  
=  $P(Z \le z - \epsilon, Z_n \le z) + P(Z \le z - \epsilon, Z_n > z)$   
 $\le F_n(z) + P(|Z_n - Z| > \epsilon)$ 

Hence,

 $F(z-\epsilon) - P(|Z-Z_n| > \epsilon) \le F_n(z) \le F(z+\epsilon) + P(|Z_n-Z| > \epsilon)$ 

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Thus,  $n \to \infty$ :

$$F(z-\epsilon) \le \liminf_{n\to\infty} F_n(z) \le \limsup_{n\to\infty} F_n(z) \le F(z+\epsilon)$$

Recall:

$$\liminf_{n \to \infty} F_n(z) \equiv \lim_{n \to \infty} \left( \inf_{m \ge n} F_m(z) \right) \le \limsup_{n \to \infty} F_n(z) \equiv \lim_{n \to \infty} \left( \sup_{m \ge n} F_m(z) \right)$$

Hence,

 $F(z-\epsilon) - P(|Z-Z_n| > \epsilon) \le F_n(z) \le F(z+\epsilon) + P(|Z_n-Z| > \epsilon)$ 

Thus,  $n \to \infty$ :

$$F(z-\epsilon) \le \liminf_{n\to\infty} F_n(z) \le \limsup_{n\to\infty} F_n(z) \le F(z+\epsilon)$$

This holds for  $\forall \epsilon > 0, \epsilon \to 0$ :  $F(z) = \lim_{n \to \infty} F_n(z)$ 

# Hoeffding's Bound

Let  $X_i$  for  $i \in \{1, ..., n\}$  be independent r.v. with  $a_i \leq X_i \leq b_i$ , and  $\overline{X}_n = \frac{1}{n} \sum_i X_i$ , then:

$$\Pr(|\overline{X}_n - \operatorname{E}[\overline{X}_n]| \ge t) \le 2 \exp\left(-\frac{2n^2 t^2}{\sum_{i=1}^n (b_i - a_i)^2}\right)$$

#### **Union Bound**

For events  $A_i$ ,  $\Pr(\cup_i A_i) \leq \sum_i \Pr(A_i)$ 

#### **Empirical Optimization Set-up**

Let  $\mathcal{H} = \{h_1, ..., h_m\}, \mathcal{D} = \{(X_1, Y_1), \dots, (X_n, Y_n)\}$ 

Suppose we wish to find the best classifier  $h \in \mathcal{H}$  to minimize its true risk  $R(h) \equiv \mathbb{E}[h(X) \neq Y].$ 

Define:

$$\hat{R}_n(h) \equiv \frac{1}{n} \sum_{i=1}^n I\{h(X_i) \neq Y_i\}$$
$$\hat{h} \equiv \operatorname{argmin}_{h \in \mathcal{H}} \hat{R}_n(h)$$
$$h^* \equiv \operatorname{argmin}_{h \in \mathcal{H}} R(h)$$

We want to bound  $|R(\hat{h}) - R(h^*)|$  with high probability. To do so, we prove that  $\sup_{h \in \mathcal{H}} |\hat{R}_n(h) - R(h)| < \epsilon$  with high probability.

$$\Pr\left(\sup_{h\in\mathcal{H}} |\hat{R}_n(h) - R(h)| < \epsilon\right)$$
$$= \Pr\left(\forall i \in \{1, \dots, m\}, |\hat{R}_n(h_i) - R(h_i)| < \epsilon\right)$$

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$$= 1 - \Pr\left(\exists i |\hat{R}_n(h_i) - R(h_i)| \ge \epsilon\right)$$

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$$\ge 1 - \sum_{i=1}^m \Pr\left(|\hat{R}_n(h_i) - R(h_i)| \ge \epsilon\right)$$

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$$= 1 - \sum_{i=1}^m \Pr\left(\left|\frac{1}{n}\sum_{j=1}^n I\{h_i(X_j) \neq Y_i\} - \mathbb{E}[I\{h_i(X) \neq Y\}]\right| \ge \epsilon\right)$$

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$$\ge 1 - 2m \exp\left(-2n\epsilon^2\right)$$

Suppose  $\sup_{h \in \mathcal{H}} |\hat{R}_n(h) - R(h)| < \epsilon$  holds.

 $R(h^*) \leq R(\hat{h})$ 

 $R(h) \equiv \mathbb{E}[h(X) \neq Y] \qquad \hat{R}_n(h) \equiv \frac{1}{n} \sum_{i=1}^n I\{h(X_i) \neq Y_i\}$  $h^* \equiv \operatorname{argmin}_{h \in \mathcal{H}} R(h) \qquad \hat{h} \equiv \operatorname{argmin}_{h \in \mathcal{H}} \hat{R}_n(h)$ 

Suppose  $\sup_{h \in \mathcal{H}} |\hat{R}_n(h) - R(h)| < \epsilon$  holds.

$$\begin{aligned} R(h^*) &\leq R(\hat{h}) \\ &\leq \hat{R}_n(\hat{h}) + \epsilon \quad \text{since } R(\hat{h}) - \hat{R}_n(\hat{h}) < \epsilon \end{aligned}$$

$$\begin{aligned} R(h) &\equiv \mathbb{E}[h(X) \neq Y] \\ h^* &\equiv \operatorname{argmin}_{h \in \mathcal{H}} R(h) \end{aligned} \quad \hat{R}_n(h) \equiv \frac{1}{n} \sum_{i=1}^n I\{h(X_i) \neq Y_i\} \\ \hat{h} &\equiv \operatorname{argmin}_{h \in \mathcal{H}} \hat{R}_n(h) \end{aligned}$$

Suppose  $\sup_{h \in \mathcal{H}} |\hat{R}_n(h) - R(h)| < \epsilon$  holds.

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$$\le R(h^*) + 2\epsilon \quad \text{since } \hat{R}_n(h^*) - R(h^*) < \epsilon$$

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Thus, 
$$|R(\hat{h}) - R(h^*)| < 2\epsilon$$

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Thus,  $|R(\hat{h}) - R(h^*)| < 2\epsilon$ , with probability at least  $1 - 2m \exp\left(-2n\epsilon^2\right)$ 

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Thus,  $|R(\hat{h}) - R(h^*)| < 2\epsilon$ , with probability at least  $1 - 2m \exp\left(-2n\epsilon^2\right)$ 

And,  $|R(\hat{h}) - R(h^*)| < 2\sqrt{\frac{\log(2m) - \log(\delta)}{2n}}$ , with probability at least  $1 - \delta$ 

#### Thank You

$$F_n(z) = P(Z_n \le z)$$
  
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Hence,

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Thus,  $n \to \infty$ :

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This holds for  $\forall \epsilon > 0, \epsilon \to 0$ :  $F(z) = \lim_{n \to \infty} F_n(z)$ 

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