

Hilbert Schmidt Independence Criterion

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1 Measuring Independence

- Covariance Operator
- Hilbert Space Methods
- A Test Statistic and its Analysis

2 Independent Component Analysis

- ICA Primer
- Examples

3 Feature Selection

- Problem Setting
- Algorithm
- Results

Measuring Independence

Problem

- Given $\{(x_1, y_1), \dots, (x_m, y_m)\} \sim \Pr(x, y)$ determine whether $\Pr(x, y) = \Pr(x) \Pr(y)$.
- Measure degree of dependence.

Applications

- Independent component analysis
- Dimensionality reduction and feature extraction
- Statistical modeling

Indirect Approach

- Perform density estimate of $\Pr(x, y)$
- Check whether the estimate approximately factorizes

Direct Approach

- Check properties of factorizing distributions
- E.g. kurtosis, covariance operators, etc.

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Covariance Operator

Linear Case

For linear functions $f(x) = w^\top x$ and $g(y) = v^\top y$ the covariance is given by

$$\text{Cov}\{f(x), g(y)\} = w^\top C v$$

This is a bilinear operator on the space of linear functions.

Nonlinear Case

Define C to be the operator with $(f, g) \rightarrow \text{Cov}\{f, g\}$.

Theorem

C is a bilinear operator in f and g .

Proof.

We only show linearity in f : $\text{Cov}\{\alpha f, g\} = \alpha \text{Cov}\{f, g\}$.

Moreover, for $f + f'$ the covariance is additive. □

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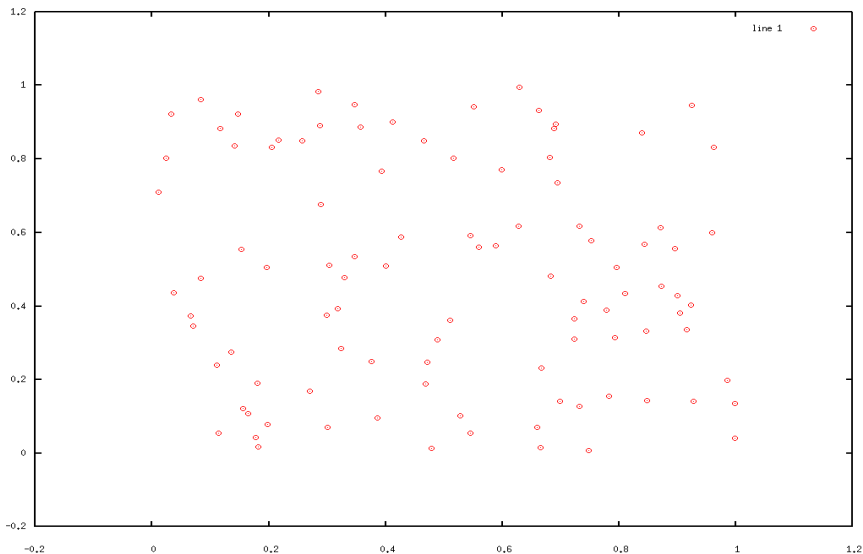
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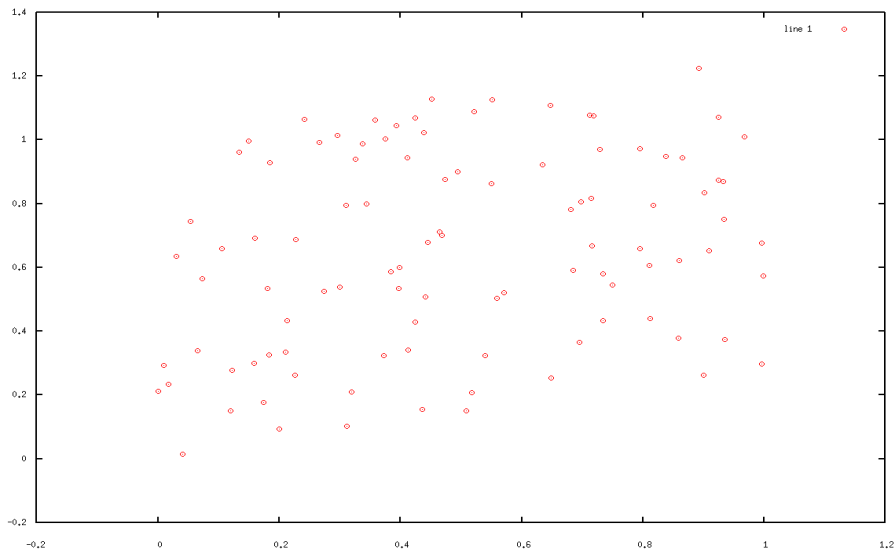
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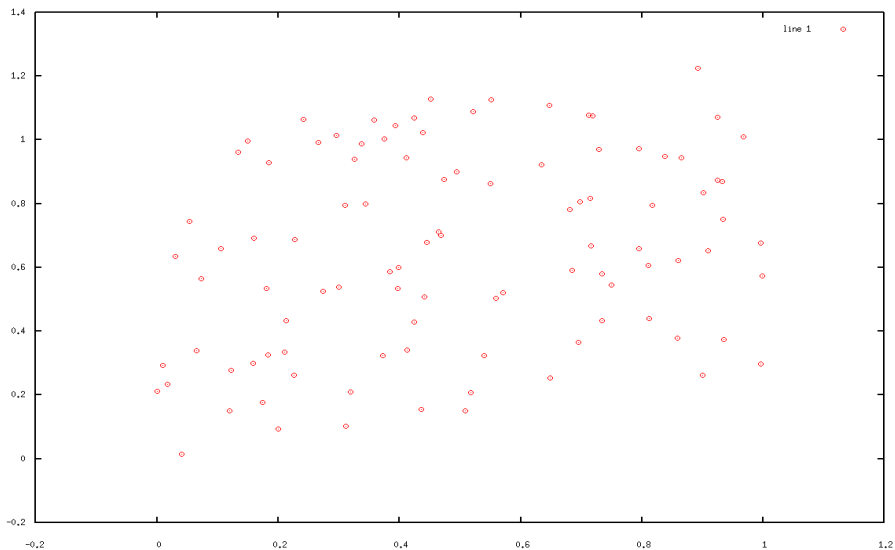
Independent random variables



Dependent random variables



Or are we just unlucky?



Covariance operators

Criterion (Renyi, 1957)

Test for independence by checking whether $C = 0$.

Reproducing Kernel Hilbert Space

- Kernels k, l on \mathcal{X}, \mathcal{Y} with associated RKHSs \mathcal{F}, \mathcal{G} .
- Assume bounded k, l on domain.

Mean operator

$$\langle \mu_x, f \rangle = \mathbf{E}_x[f(x)] \text{ and } \langle \mu_y, g \rangle = \mathbf{E}_y[g(y)]$$

Covariance operator

Define covariance operator C via bilinear form

$$f^\top C_{xy} g = \text{Cov}\{f, g\} = \mathbf{E}_{x,y}[f(x)g(y)] - \mathbf{E}_x[f(x)] \mathbf{E}_y[g(y)]$$

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Hilbert Space Representation

Theorem

Provided that k, l are universal kernels $\|C_{xy}\| = 0$ if and only if x, y are independent.

Proof.

Step 1: If x, y are dependent then there exist some $[0, 1]$ -bounded range f^*, g^* with $\text{Cov}\{f^*, g^*\} = \epsilon > 0$.

Step 2: Since k, l are universal there exist ϵ' approximation of f^*, g^* in \mathcal{F}, \mathcal{G} such that covariance of approximation does not vanish.

Step 3: Hence the covariance operator C_{xy} is nonzero.



A Test Statistic

Covariance operator

$$C_{xy} = \mathbf{E}_{x,y} [k(x, \cdot)l(y, \cdot)] - \mathbf{E}_x [k(x, \cdot)] \mathbf{E}_y [l(y, \cdot)]$$

Operator Norm

Use the norm of C_{xy} to test whether x and y are independent. It also gives us a **measure** of dependence.

$$\text{HSIC}(\Pr, \mathcal{F}, \mathcal{G}) := \|C_{xy}\|^2$$

where $\|\cdot\|$ denotes the Hilbert-Schmidt norm.

Frobenius Norm

For matrices we can define

$$\|M\|^2 = \sum_{ij} M_{ij}^2 = \text{tr } M^T M.$$

Hilbert-Schmidt norm is generalization of Frobenius norm.

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Computing $\|C_{xy}\|^2$

Rank-one operators

For rank-one terms we have

$$\|f \otimes g\|^2 = \langle f \otimes g, f \otimes g \rangle_{HS} = \|f\|^2 \|g\|^2.$$

Joint expectation

By construction of C_{xy} we exploit linearity and obtain

$$\begin{aligned} \|C_{xy}\|^2 &= \langle C_{xy}, C_{xy} \rangle_{HS} \\ &= \{ \mathbf{E}_{x,y} \mathbf{E}_{x',y'} - 2\mathbf{E}_{x,y} \mathbf{E}_{x'} \mathbf{E}_{y'} + \mathbf{E}_x \mathbf{E}_y \mathbf{E}_{x'} \mathbf{E}_{y'} \} \\ &\quad \left[\langle k(x, \cdot) l(y, \cdot), k(x', \cdot) l(y', \cdot) \rangle_{HS} \right] \\ &= \{ \mathbf{E}_{x,y} \mathbf{E}_{x',y'} - 2\mathbf{E}_{x,y} \mathbf{E}_{x'} \mathbf{E}_{y'} + \mathbf{E}_x \mathbf{E}_y \mathbf{E}_{x'} \mathbf{E}_{y'} \} \\ &\quad [k(x, x') l(y, y')] \end{aligned}$$

This is well-defined if k, l are bounded.

Estimating $\|C_{xy}^2\|$

Empirical criterion

$$\text{HSIC}(Z, \mathcal{F}, \mathcal{G}) := \frac{1}{(m-1)^2} \text{tr} KHLH$$

where $K_{ij} = k(x_i, x_j)$, $L_{ij} = l(y_i, y_j)$ and $H_{ij} = \delta_{ij} - m^{-2}$.

Theorem

$$\mathbf{E}_Z [\text{HSIC}(Z, \mathcal{F}, \mathcal{G})] = \text{HSIC}(\Pr_{xy}, \mathcal{F}, \mathcal{G}) + O(1/m)$$

Proof: Sketch only.

Expand $\text{tr} KHLH$ into terms of pairs, triples and quadruples of indices of non-repeated terms, which lead to the proper expectations and bound the rest by $O(m^{-1})$. □

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Uniform convergence bounds for $\|C_{xy}^2\|$

Theorem (Recall Hoeffding's theorem for U Statistics)

For averages over functions on r variables

$$u := \frac{1}{\binom{m}{r}} \sum_{i_r^m} g(x_{i_1}, \dots, x_{i_r})$$

which are bounded by $a \leq u \leq b$ we have

$$\Pr_u \{u - \mathbf{E}_u[u] \geq t\} \leq \exp\left(-\frac{2t^2 \lceil m/r \rceil}{(b-a)^2}\right)$$

In our statistic we have terms of 2, 3, and 4 random variables.

Uniform convergence bounds for $\|C_{xy}^2\|$

Corollary

Assume that $k, l \leq \cdot$. Then at least with probability $1 - \delta$

$$\left| \text{HSIC}(Z, \mathcal{F}, \mathcal{G}) - \text{HSIC}(\mathbf{Pr}_{xy}, \mathcal{F}, \mathcal{G}) \right| \leq \sqrt{\frac{\log 6/\delta}{0.24m}} + \frac{C}{m}$$

Proof.

Bound each of the three terms separately via Hoeffding's theorem. □

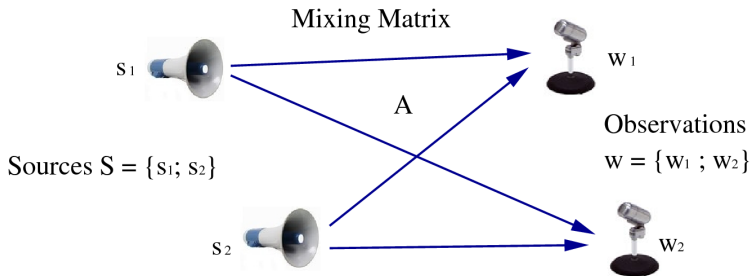
- 1 Measuring Independence
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Blind Source Separation

Data

$w = Ms$, where all s_i are mutually **independent**.

The Cocktail Party Problem



Task

Recover the sources S and mixing matrix M given W .

Independent Component Analysis

Whitening

Rotate, center, and whiten data before separation. This is always possible.

Optimization

- We cannot recover scale of data anyway.
- Need to find **orthogonal** matrix U such that $Uw = s$ leads to independent random variables.
- Optimization on the **Stiefel manifold**.
- Could do this by a Newton method.

Important Trick

- Kernel matrix could be **huge**.
- Use reduced-rank representation. We get

$$\text{tr } H(AA^\top)H(BB^\top) = \|A^\top HB\|^2 \text{ instead of } \text{tr } HKHL.$$

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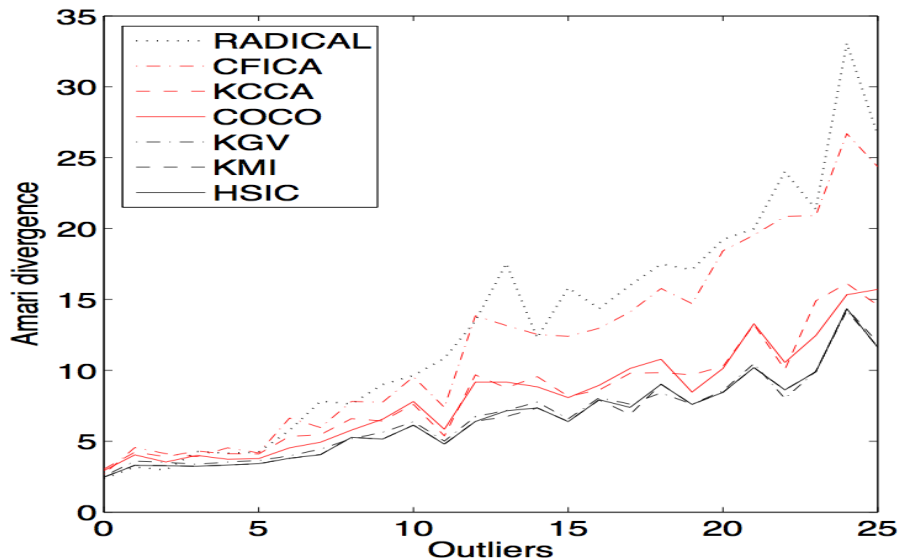
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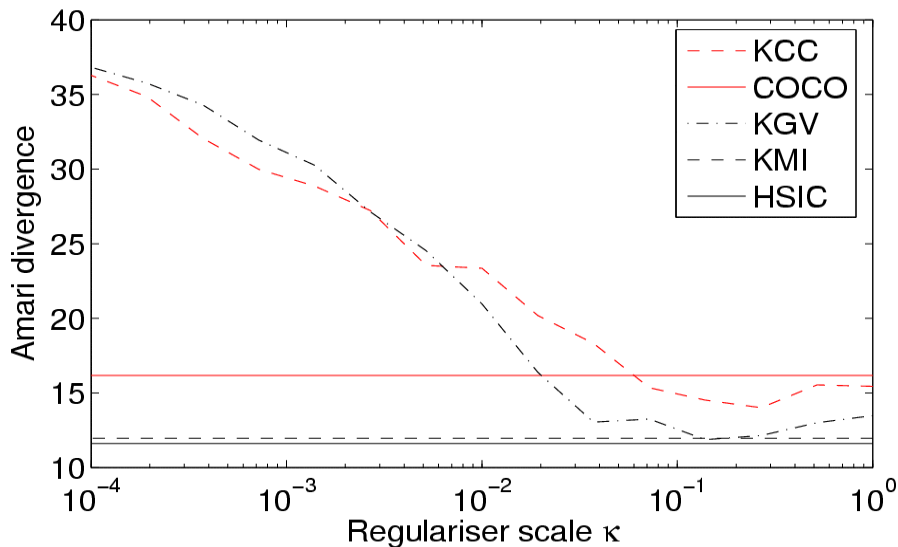
ICA Experiments

n	m	Rep.	FICA	Jade	IMAX	RAD	CFIC	KCC	COg	COI	KGV	KMIg	KMII	HSICg	HSICI
2	250	1000	10.5± 0.4	9.5 ± 0.4	44.4± 0.9	5.4 ± 0.2	7.2 ± 0.3	7.0 ± 0.3	7.8 ± 0.3	7.0 ± 0.3	5.3 ± 0.2	6.0 ± 0.2	5.7 ± 0.2	5.9 ± 0.2	5.8 ± 0.3
2	1000	1000	6.0 ± 0.3	5.1 ± 0.2	11.3± 0.6	2.4 ± 0.1	3.2 ± 0.1	3.3 ± 0.1	3.5 ± 0.1	2.9 ± 0.1	2.3 ± 0.1	2.6 ± 0.1	2.3 ± 0.1	2.6 ± 0.1	2.4 ± 0.1
4	1000	100	5.7 ± 0.4	5.6 ± 0.4	13.3± 1.1	2.5 ± 0.1	3.3 ± 0.2	4.5 ± 0.4	4.2 ± 0.3	4.6 ± 0.6	3.1 ± 0.6	4.0 ± 0.7	3.5 ± 0.7	2.7 ± 0.1	2.5 ± 0.2
4	4000	100	3.1 ± 0.2	2.3 ± 0.1	5.9 ± 0.7	1.3 ± 0.1	1.5 ± 0.1	2.4 ± 0.5	1.9 ± 0.1	1.6 ± 0.1	1.4 ± 0.1	1.4 ± 0.05	1.2 ± 0.05	1.3 ± 0.05	1.2 ± 0.05
8	2000	50	4.1 ± 0.2	3.6 ± 0.2	9.3 ± 0.9	1.8 ± 0.1	2.4 ± 0.1	4.8 ± 0.9	3.7 ± 0.9	5.2 ± 1.3	2.6 ± 0.3	2.1 ± 0.1	1.9 ± 0.1	1.9 ± 0.1	1.8 ± 0.1
8	4000	50	3.2 ± 0.2	2.7 ± 0.1	6.4 ± 0.9	1.3 ± 0.05	1.6 ± 0.1	2.1 ± 0.2	2.0 ± 0.1	1.9 ± 0.1	1.7 ± 0.2	1.4 ± 0.1	1.3 ± 0.05	1.4 ± 0.05	1.3 ± 0.05
16	5000	25	2.9 ± 0.1	3.1 ± 0.3	9.4 ± 1.1	1.2 ± 0.05	1.7 ± 0.1	3.7 ± 0.6	2.4 ± 0.1	2.6 ± 0.2	1.7 ± 0.1	1.5 ± 0.1	1.5 ± 0.1	1.3 ± 0.05	1.3 ± 0.05

Outlier Robustness



Automatic Regularization



Linear mixture of independent sources

- Remove mean and whiten for preprocessing
- Use HSIC as measure of dependence
- Find best rotation to demix the data

Performance

- HSIC is very robust to outliers
- General purpose criterion
- Best performing algorithm (Radical) is designed for *linear* ICA, HSIC is a *general purpose criterion*
- Low rank decomposition makes optimization feasible

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Feature Selection

The Problem

- Large number of features
- Select a small subset of them

Basic Idea

- Find features such that the distributions $p(x|y = 1)$ and $p(x|y = -1)$ are as different as possible.
- Use a two-sample test for that.

Important Tweak

We can find a similar criterion to measure dependence between data and labels (by computing the Hilbert-Schmidt norm of covariance operator).

Recursive Feature Elimination

Algorithm

- Start with full set of features
- Adjust kernel width to pick up maximum discrepancy
- Find feature which decreases dissimilarity the least
- Remove this feature
- Repeat

Applications

- Binary classification (standard MMD criterion)
- Multiclass
- Regression

Algorithm 1 Feature Selection via Backward Elimination

Input: The full set of features \mathcal{S}

Output: An ordered set of features \mathcal{S}^\dagger

1: $\mathcal{S}^\dagger \leftarrow \emptyset$

2: **repeat**

3: $\sigma_0 \leftarrow \arg \max_{\sigma} \text{HSIC}(\sigma, \mathcal{S})$

4: $i \leftarrow \arg \max_i \text{HISC}(\sigma_0, \mathcal{S} \setminus \{i\}), i \in \mathcal{S}$

5: $\mathcal{S} \leftarrow \mathcal{S} \setminus \{i\}$

6: $\mathcal{S}^\dagger \leftarrow \mathcal{S}^\dagger \cup \{i\}$

7: **until** $\mathcal{S} = \emptyset$

Comparison to other feature selectors

Synthetic Data

Table 1: Classification error (%) after selecting features using BAHSIC and 6 other methods.

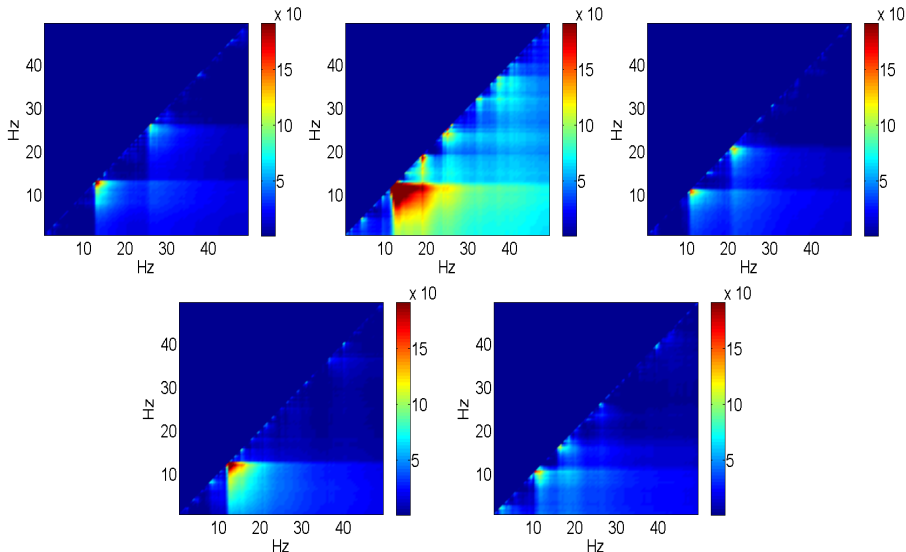
Method	Fisher	FSV	L0	MI	R2W2	RFE	BAHSIC
WL-6	10.0±4.5	2.0±2.0	0.0±0.0	6.0±3.1	0.0±0.0	0.0±0.0	0.0±0.0
WN-2	57.0±3.7	58.0±5.3	2.0±1.3	18.0±2.9	54.0±6.5	2.0±1.3	1.0±1.0

Brain Computer Interface Data

Table 2: Classification errors (%) on BCI data after selecting a frequency range.

Subject	aa	al	av	aw	ay
CSP(8-40Hz)	17.5±2.5	3.1±1.2	32.1±2.5	7.3±2.7	6.0±1.6
CSSP	14.9±2.9	2.4±1.3	33.0±2.7	5.4±1.9	6.2±1.5
CSSSP	12.2±2.1	2.2±0.9	31.8±2.8	6.3±1.8	12.7±2.0
BAHSIC	13.7±4.3	1.9±1.3	30.5±3.3	6.1±3.8	9.0±6.0

Frequency Band Selection



Microarray Feature Selection

Goal

- Obtain small subset of features for estimation
- Reproducible feature selection

Results

Table 3: Comparison between SVM-RFE and BAHSIC for bioinformatics data. From top to bottom: data set description, classification errors in (%), and feature stability.

Dataset	Lymphoma		Yeast		Colon	Berchuck
Dim	4026		79		2000	22283
Sample	42/11/9		121/35/27/14/11		40/22	30/24
	MC	OVR	MC	OVR		
SVM	32.4±6.9	32.4±6.9	5.3±2.1	5.8±1.8	17.6±5.1	43.3±6.9
RFE	12.86±3.30	0.00±0.00	30.36±2.39	6.76±2.10	22.38±6.05	30.00±7.57
BAHSIC	0.00±0.00	0.00±0.00	5.79 ±1.99	4.81±1.59	15.71±5.27	19.33±6.30
RFE	0.77±0.09	0.46±0.28	0.41±0.31	0.39±0.32	0.38±0.11	0.57±0.28
BAHSIC	0.96±0.03	0.96±0.03	0.82±0.14	0.82±0.14	0.90±0.06	0.73±0.19

Table 4: Root mean square error (RMSE) of support vector regression with and without HSIC

Method	Sample	Dim	Feature	ϵ -SVR	RAND	BAHSIC
Pyrim	55	27	5	0.112±0.067	0.092±0.073	0.085±0.066
Triaz	186	60	2	0.147±0.027	0.157±0.036	0.144±0.033
Bodyfat	227	14	7	0.0019±0.0026	0.0019±0.0026	0.0019±0.0024

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Shameless Plugs

Looking for a job ... talk to me!

- Alex.Smola@nicta.com.au (<http://www.nicta.com.au>)

Positions

- PhD scholarships
- Postdoctoral positions, Senior researchers
- Long-term visitors (sabbaticals etc.)

More details on kernels

- <http://sml.nicta.com.au>
- <http://www.kernel-machines.org>
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