

# **Exponential Families** Classification and Novelty Detection

S.V.N. "Vishy" Vishwanathan

vishy@axiom.anu.edu.au

#### National ICT of Australia and Australian National University

Thanks to Alex Smola, Thomas Hofmann and Stéphane Canu

# **Overview**



- Review of Exponential Family
- Log Partition Function
- Maximum Likelihood Estimation
- MAP Estimation
- Conditional Densities
- Gaussian Processes and the Normal Prior
- Novelty Detection
- Large Margin Classifiers



## **Basic Equation:**

We will model densities by

$$p(\mathbf{x}; \theta) = \exp(\langle \phi(\mathbf{x}), \theta \rangle - g(\theta))$$

## Why Exponential Families:

- Dense in space of densities
- Interpretent the sector  $\phi(\mathbf{x})$  closely related to kernels
- Solution State State
- Conditional models are easy to derive
- Close connections to graphical models

#### Where is the Catch:

- Statisticians work with explicitly parameterized  $\phi(\mathbf{x})$
- The log-partition function is difficult to compute



## **Basic Equation:**

Some algebra gives us

$$g(\theta) = \log \int_{\mathcal{X}} \exp(\langle \phi(\mathbf{x}), \theta \rangle) \, d\, \mathbf{x}$$

**\checkmark** Computing this integral over  $\mathcal{X}$  is painful

# **Moment Generating Function:**

Derivatives of  $g(\theta)$  generate moments of  $\phi(\mathbf{x})$ 

$$\partial_{\theta} g(\theta) = \mathbb{E}_{p(\mathbf{x};\theta)} [\phi(\mathbf{x})]$$
  
$$\partial_{\theta}^{2} g(\theta) = \operatorname{Var}_{p(\mathbf{x};\theta)} [\phi(\mathbf{x})]$$

#### **Other Properties:**

The log-partition function is convex

It is extremely smooth and differentiable ( $C^{\infty}$  function)

# **Universal Density Estimators**



# Setting:

- Let X be a measurable set and  $k: X \times X \to \mathbb{R}$  a kernel
- $\textbf{ Let } f(\cdot) = \langle \phi(\cdot), \theta \rangle_{\mathcal{H}} \text{ and } f(\mathbf{x}) = \langle f(\cdot), k(\cdot, \mathbf{x}) \rangle_{\mathcal{H}}$
- In the set of continuous and bounded densities on X be  $\mathcal{P}$
- Furthermore let  $\mathcal{H}$  be dense in  $C^0(\mathfrak{X})$

#### **Universal Density Estimators:**

The densities  $p_f(\mathbf{x}) := \exp(f(\mathbf{x}) - g_f(\theta))$  are dense in 𝒫

#### **Proof Sketch:**

- Find a  $f(\mathbf{x})$  close to given  $\bar{\mathbf{p}}(\mathbf{x})$
- **Show that**  $\int_{\mathfrak{X}} \exp(f(\mathbf{x})) d\mathbf{x}$  is bounded
- $\checkmark$  It follows that  $|\log p_f(\mathbf{x}) \log \bar{\mathbf{p}}(\mathbf{x})|$  is small
- Solution Hence conclude that  $|p_f(\mathbf{x}) \bar{\mathbf{p}}(\mathbf{x})|$  is small



# Why Condition:

- We are given  $(\mathbf{x}_i, y_i)$  pairs
- $\checkmark$  Given a new data point we want to predict its label y
- $\checkmark$  We don't want to waste modeling effort on  $\mathbf{x}$

#### The Answer:

By Bayes rule we know

$$p(y|\mathbf{x}; \theta) = \frac{p(\mathbf{x}, y; \theta)}{p(\mathbf{x}; \theta)}$$

#### **The Exponential Family:**

If  $p(y|\mathbf{x}; \theta)$  is a member of the exponential family  $p(y|\mathbf{x}; \theta) = \exp(\langle \phi(\mathbf{x}, y), \theta \rangle - g(\theta | \mathbf{x}))$   $g(\theta | \mathbf{x}) = \log \int_{\mathcal{Y}} \exp(\langle \phi(\mathbf{x}, \bar{\mathbf{y}}), \theta \rangle d \bar{\mathbf{y}}$ 



#### **Basic Idea:**

- Data drawn from a conditional exponential density
- Find the  $\theta$  which maximizes  $p(y | \mathbf{X}; \theta)$

#### The Model:

By iid assumption

$$\log p(y | \mathbf{X}; \theta) = \sum_{i=1}^{m} \log p(y_i | \mathbf{x}_i; \theta)$$

#### The Solution:

**9** By setting  $\partial_{\theta} p(y | \mathbf{X}; \theta) = 0$  we get

$$\mathbb{E}_{p(y|\mathbf{x};\theta)}[\phi(\mathbf{x},y)] = \frac{1}{m} \sum_{i=1}^{m} \phi(\mathbf{x}_i,y_i)$$



#### **Basic Idea:**

- **J** We assume that  $\theta$  is a random variable
- Also assume a prior (belief) over  $\theta$
- Now the data updates our belief about the prior

#### **The Normal Prior:**

- $\checkmark$  We assume  $\theta \sim \mathcal{N}(0, \sigma^2)$
- By Bayes rule

$$p(\theta | \mathbf{X}, y) \propto p(y | \mathbf{X}; \theta) p(\theta)$$

#### **The Solution:**

• By setting 
$$\partial_{\theta} - \log p(\theta | \mathbf{X}, y) = 0$$
 we get  
 $\mathbb{E}_{p(y | \mathbf{x}; \theta)}[\phi(\mathbf{x}, y)] = \frac{1}{m} \sum \phi(\mathbf{x}_i, y_i) - \frac{\theta}{m\sigma^2}$ 

# Key Idea:

- **9** Let  $t: \mathfrak{X} \to \mathbb{R}$  be a stochastic process
- Fix any  $\{\mathbf{x}_1, \ldots, \mathbf{x}_m\}$
- For a GP  $\{t(\mathbf{x}_1), \ldots, t(\mathbf{x}_m)\}$  are jointly normal

#### Parameters of a GP:

Mean

$$\mu(\mathbf{x}) := \mathbb{E}[t(\mathbf{x})]$$

Covariance function (kernel)

$$k(\mathbf{x}, \mathbf{x'}) := \operatorname{Cov}(t(\mathbf{x}), t(\mathbf{x'}))$$

# **Simplifying Assumption:**

 $\checkmark$  We know the form of  $k(\mathbf{x},\mathbf{x}')$ 





## Key Idea:

- $\textbf{Let } \boldsymbol{\theta} \sim \mathcal{N}(0, \sigma^2)$
- $\checkmark$  Then  $\log p(y | \mathbf{x}; \theta) + g(\theta | \mathbf{x})$  is a GP

Why?:

- $\textbf{ Observe that } \log p(y|\mathbf{x};\theta) + g(\theta|\mathbf{x}) = \langle \phi(\mathbf{x},y),\theta \rangle$
- Hence it is normally distributed
- $\checkmark$  The mean  $\mathbb{E}_{\theta}[\langle \phi(\mathbf{x},y),\theta \rangle = 0$
- The covariance is given by

$$k((\mathbf{x},y),(\mathbf{x}',y)) = \sigma^2 \langle \phi(\mathbf{x},y), \phi(\mathbf{x}',y') \rangle$$

#### **Observations:**

- $\checkmark$  Kernel can depend on both  $\mathbf x$  and y
- Extensions to multi-class problems possible
- $\checkmark$  If y has structure we can exploit it



## **Optimization Problem:**

The MAP estimate solves

$$\operatorname{argmin}_{\theta} \frac{1}{2\sigma^2} ||\theta||^2 - \sum_{i=1}^m \langle \phi(\mathbf{x}_i, y_i), \theta \rangle + g(\theta | \mathbf{x}_i)$$

By the representer theorem

$$\theta = \sum_{i=1}^{m} \sum_{y \in \mathcal{Y}} \alpha_{iy} \phi(\mathbf{x}_i, y_i)$$

#### **Observations:**

- Convex Optimization problem
- **I**f  $|\mathcal{Y}|$  is large we are in trouble
- In case of binary classification we use  $\phi(\mathbf{x}, y) = y\phi(\mathbf{x})$



# Key Idea:

- $\checkmark$  We estimate  $p(\mathbf{x} | \theta)$  based on  $\{\mathbf{x}_i\}$
- All  $\mathbf{x}_i$  with  $p(\mathbf{x}_i | \theta) < p_0$  are novel

# **Tightening the Belt:**

- Don't waste modeling effort on high density regions
- Only shape of  $p(\mathbf{x} | \theta)$  is important

# The Solution:

Estimate

$$\min\left(\frac{p(\mathbf{x}_i | \theta)}{p_0}, 1\right)$$

• We use 
$$p_0 = \exp(\rho - g(\theta))$$

Jelps get rid of pesky  $g(\theta)$  term



# **Exponential Family:**

Using the iid assumption our objective function is

$$\operatorname{argmax}_{\theta} \prod_{i=1}^{m} \min\left(\frac{p(\mathbf{x}_i \mid \theta)}{p_0}, 1\right) p(\theta)$$

## The Final Form:

If we assume a normal prior and use log likelihoods

$$\operatorname{argmin}_{\theta} \sum_{i=1}^{m} \max(\rho - \sum_{j} k(\mathbf{x}_{i}, \mathbf{x}_{j}, 0) + \frac{1}{2^{\sigma^{2}}} ||\theta||^{2}$$

**Solution** Exactly the problem solved by the Single class SVM! **The**  $\nu$ **-Trick:** 

# **Odds Ratio**

# **Basic Idea:**

- In OCR classification 0 and 8 are frequently confused
- Digits like 0 and 1 are generally well classified
- Worst confused class is a measure of margin

# The Solution:

If we consider the ratio

$$R(\mathbf{x}, y, \theta) = \min_{y \neq y'} \exp(\langle \phi(\mathbf{x}, y) - \phi(\mathbf{x}, y'), \theta \rangle)$$

- Measure of confusion with the next best class
- We can interpret this as the margin

#### The Consequences:

- SVM like large margin classifiers are special cases
- Extensions to multi-class setting natural

# **Algorithm:**

- Arguably the simplest algorithm in machine learning!
- Maintains a weight vector  $\theta$
- **9** Given  $(\mathbf{x}_i, y_i)$ 
  - If  $y_i \langle \phi(\mathbf{x}_i), \theta \rangle \geq 0$  do nothing
  - Else  $\theta = \theta + y_i \phi(\mathbf{x}_i)$

• Notice how 
$$\theta = \sum_j y_j \phi(\mathbf{x}_j)$$

## Novikoff's Theorem:

- **9** Given  $S = \{(\mathbf{x}_i, y_i\}$  non-trivial
- Let  $R = \max_i ||\phi(\mathbf{x}_i)||$  be the radius of the samples
- $\exists \theta^* \text{ such that } ||\theta^*|| = 1 \text{ and } y_i \langle \theta^*, \phi(\mathbf{x}_i) \rangle \ge \gamma > 0$
- Interpretation of the second seco



#### The Question:

- **•** Consider  $\mathcal{Y} = \{1, 2, 3\}$
- **Suppose we know**  $p(1|\mathbf{x}_i) = 0.4$
- **Solution** Can we conclude that  $\mathbf{y}_i = 1$ ?

## The Answer:

- ▶ No! because we might have  $p(2|\mathbf{x}_i) = 0.5$
- We want the true class probability to be peaked

#### The Consequences:

- This is equivalent to maximizing the log odds ratio
- Using a normal prior on  $\theta$  we can solve

$$\min \frac{1}{2} ||\theta||^2$$
  
s.t.  $\log R(\mathbf{x}_i, y_i, \theta) \ge 1$ 

CT AUSTRAL



# **Recovering SVM's**

# The Problem:



 $\min \frac{1}{2} ||\theta||^2$ s.t.  $\log R(\mathbf{x}_i, y_i, \theta) \ge 1$ 

Recall that

$$\log R(\mathbf{x}, y, \theta) = \max_{y \neq y'} \langle \phi(\mathbf{x}, y) - \phi(\mathbf{x}, y'), \theta \rangle$$

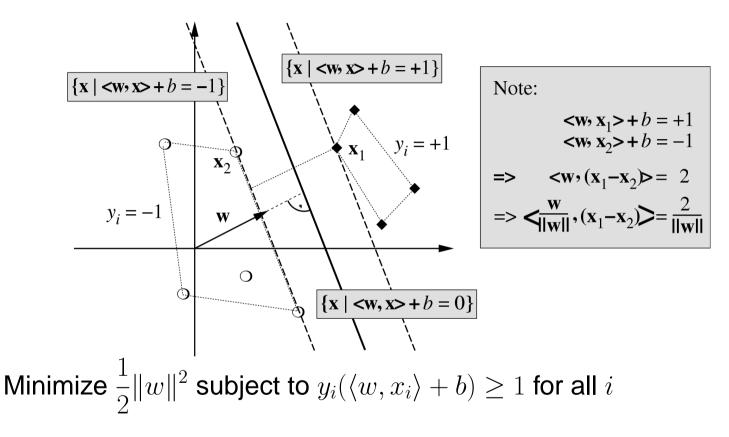
#### The Binary Case:

- Again recall that  $\phi(\mathbf{x},y) = y\phi(\mathbf{x})$  in the binary case
- $\textbf{ Hence } R(\mathbf{x},y,\theta) = 2 \langle \phi(\mathbf{x}),\theta \rangle$
- The equivalent optimization problem is

$$\min \frac{1}{2} ||\theta||^2$$
  
s.t. $y_i \langle \phi(\mathbf{x}_i), \theta \rangle \ge 1$ 



# **Optimal Separating Hyperplane**



#### 

#### **Slack Variables:**

- Data might not be linearly separable in feature space
- To avoid over fitting ignore noisy points
- We modify the optimization problem

$$\min \frac{1}{2} ||\theta||^2 + C \sum_i \xi_i$$
  
s.t. $R(\mathbf{x}_i, y_i, \theta) \ge 1 - \xi_i \quad \xi_i \ge 0$ 

# **Upper Bound on Error:**

If we define

$$\xi_i(\theta) = \max\{0, 1 - R(\mathbf{x}_i, y_i, \theta)\}$$

then

$$\frac{1}{m}\sum_{i=1}^{m}\xi_i(\theta) \ge \frac{1}{m}\sum_{i=1}^{m}\delta(y_i, \operatorname{sign}(\log R(\mathbf{x}_i, y_i, \theta)))$$

# **Slack Variables:**

- We include a slack term for every linear constraint
- The optimization problem becomes

$$\min \frac{1}{2} ||\theta||^2 + C \sum_i \sum_{y \neq y_i} \xi_{iy}$$
  
s.t. $\langle \phi(\mathbf{x}_i, y_i) - \phi(\mathbf{x}_i, y), \theta \rangle \ge 1 - \xi_{iy} \quad \xi_{iy} \ge 0$ 

# **Upper Bound on Ranking Error:**

Now we can write a bound

$$\frac{1}{m}\sum_{i=1}^{m}\xi_{iy}(\theta) \ge \frac{1}{m}\sum_{i=1}^{m}|\{y \neq y_i : \langle \phi(\mathbf{x}_i, y), \theta \rangle \ge \langle \phi(\mathbf{x}_i, y_i), \theta \rangle\}|$$

#### **Comments:**

- $\checkmark$  More constraints  $\implies$  harder problem to solve
- Solution might not be sparse!

# **Questions?**

